

# Farmers' Willingness to Pay for Improved Soil Conservation Practices in Kuyu Woreda, North Shoa Zone of Oromia, Ethiopia: Application of Contingent Valuation Method

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## Abstract

Land degradation particularly soil erosion is one of the most severe environmental problems in Ethiopia. Achieving sustainable pathways out of land degradation problem and poverty require active participation of people in conservation practices and understand how farmers value the soil conservation. The objective of this study is to investigate the farmers' willingness to pay for improved soil conservation practices on communal land in Kuyu woreda, North Shoa, Oromia Regional State of Ethiopia. Double-bounded dichotomous format of contingent valuation method is used to elicit respondents' WTP. Data collected from 184 households were used for analysis. Both the Probit and Bivariate Probit models were estimated. The result of the study indicated that most of the respondents have perceived soil erosion problem and are willing to pay for conservation practices. Factors like sex, education, livestock, income, slope of land, perception, and the frequency of extension visit have a positive and significant influence on the probability of WTP for soil conservation. The study also shows that the mean WTP estimated from the Double Bounded Dichotomous Choice formats was 85.36 Labor days per annual. The total aggregate value of soil conservation in the study is 10,909,349 person days per annum (196,368,282 ETB) from double bound CVM. The study concluded that farmers have perceived soil erosion problem and they are willing to pay in terms of labor contribution for improved soil conservation practice. Therefore, policy implication suggested was any policy intervention designed to address soil erosion problems have needed to take in to account these characteristics.

**Keywords:** Willingness to pay, Double Bounded Dichotomous Choice, soil erosion, Bivariate Probit models

## Introduction

### *Background and Rationale of the Study*

Land is a vital resource for producing food, preserving biodiversity, facilitating the natural management of water systems and acting as a carbon store; and appropriate land management can protect and maximize these services for society. Conversely, desertification, land degradation and drought (DLDD) have accelerated during the 20th and 21st centuries. The productivity of land is diminishing at an alarming pace and every year, 5-7 million hectare of agricultural land is lost worldwide. Higher input may offset lower productivity in the industrial countries, but most developing countries like

Ethiopia lack the capabilities for this. Although the problem of soil erosion is as old as settled agriculture, its extent and impact on human welfare and global environment are more now than ever before (UNCCD, 2013).

Since the 1970s soil erosion has been identified as the most severe environmental problem in the highlands of East African countries including Ethiopia (Gebremedhin, 1998). In Ethiopia, New land conservations were started in soil erosion and food deficit areas through food-for-work in the 1970 and 1980s (Shiferaw and Holden, 1998). However, the interventions needed improvement to maximize the benefit and balance current and future generation's need (Shimeles, 2012). Currently, even if Ethiopia government is undertaking community participatory measures to conserve land, still now there is the need to improvement on mobilizing local people in conservation.

In Ethiopia, particularly in Oromia regional state, soil erosion is one of the most serious environmental problems especially on the farm and communal land. Even if in the county efforts to curb the soil erosion problem are always in action, inappropriate use of land and absence of active local people mobilization are the main cause of soil erosion (OADB, 2011). Particularly, soil erosion in Kuyu Woreda is severely affecting communal and cultivation lands; and from the total area of the kuyu Woreda around 38570.546 hectares are communally owned lands. Unrestricted access (free grazing), traditional land use, and lack of awareness, absence of appropriate sustainable soil conservation measures and mainly absence of active participation of community are the main cause for soil erosion in the woreda (KWADB, 2011). Thus this environmental problem has impact on the well-being of the community in the study area.

Therefore, the challenge is how to overcome environmental degradation that affects the livelihood of the people in the country in general and Kuyu woreda in particular. Soil conservation is one of the strategies. However, achieving sustainable pathways out of the problem of Soil erosion and poverty through soil conservation strategy requires active participation of farmers and understand how farmers value the soil conservation of activities especially on communal areas of the study area. Thus, to overcome the soil erosion through understanding how farmers value the soil conservation, the non-market valuation methods have become crucial. In this study stated contingent valuation method was employed to estimate the value of improved soil conservation program in a hypothetical market.



**Soil erosion picture in Kuyu Woreda**

To our knowledge, there has no empirical study on soil conservation that applied valuation techniques on communal lands in Ethiopia in general and the study area in particular. Hence, this kind of research is helps the government and any concerned agent for intervention mechanisms and mobilize the farmers for conservation. The study is limited to examining farmers' WTP for improved soil conservation practices in Kuyu woreda, North Shoa Zone of Oromia, Ethiopia.

### ***General Objective***

The main objective of this study is to investigate the farmers' willingness to pay for improved soil conservation practices on communal lands of Kuyu woreda.

### ***Specific Objectives***

- To identify the determinants of farmers' WTP for improved soil conservation practices

- To estimate the amount of farmers willingness to pay for improved soil conservation practice
- To estimate the welfare gain using CVM
- To suggest policy implications based on the findings of the study

## **Research Methodology**

### ***Source of Data and Data Collection Method***

The primary data were collected from 200 randomly selected households from 5 Kebeles of Kuyu Woreda through CV survey questionnaire by employing face-to-face data collection techniques. A double-bounded elicitation format that improve statistical efficiency over single-bounded is used to elicit respondents' WTP in terms of labor per day (Haab and McConnell, 2002; Hanemann, 1991). In addition, the socioeconomic characteristics information of the households was collected. Secondary data collected from books, journals, articles, and so on.

### ***Population, Sample Size and Sample Techniques***

Based on the existence of communal land and erosion severity five kebeles were purposely selected from total 23 kebele of kuyu woreda. Next, two stage sampling technique was employed to select interviewed households. Then, five development teams were randomly selected from each kebele. Finally, eight households were randomly selected from each development teams by using simple random sampling technique. Hence the total number of households was 200.

### ***Survey Questionnaires Design and Elicit Format and Field work Procedure***

CV survey questionnaires have two parts. First, provide general information and try to collect information on demographic structure and socio economic level of respondents; second provides contingent valuation scenarios and elicits farmers' WTP for improved soil conservation questions. Before the final survey is implemented the pilot survey that included 20 households was carried out by using open-ended election format. The pilot survey was averaged at 50 person day per year; and based on this results three starting point price were introduced. The field survey was successfully completed with low invalid responses (8.7% protest zeros).

### ***Method of Data Analysis***

To analyses data both Descriptive analysis and Econometrics analysis were employed. Both the probit and bivariate probit models were applied for the estimation of CVM. Bivariate model in estimating the double bound CV can lead to efficiency (Geniush et al., 2005). Cameron (1994) cited in Haab (1997) propose the use of a bivariate probit contingent valuation model when respondents are offered a follow-up bid to an initial contingent valuation question. The bivariate probit CVM solves distortion which is introduced from the follow up questions of the double-bounded contingent valuation survey.

### ***Model specification***

Farmers' decision to pay or not to pay for at any time and can be modeled in a utility framework following Hanemann (1984) as:  $U_i = U_i(L, Z, q)$ . Where  $U_i$  is the utility of the household  $i$ ,  $L$  is total labour endowment of the household in a year,  $Z$  are socioeconomic characteristics of the household, whereas  $q$  is soil conservation quality as perceived by the farmer. Furthermore, let us assume that  $q^*$  as the quality after the soil conservation practice and  $q$  as the quality before the soil conservation practices is undertaken. Then:  $U_i^1(L - bid, Z, q^*) + e_i \geq U_i^0(L - bid, Z, q) + e_0$ . BID, is the initial labor payment per year,  $e_i$  and  $e_0$  are the error terms which are with zero means and independently distributed. Therefore, the probability that a household will decide to pay for the soil conservation is the probability that the conditional indirect utility function for the proposed intervention is greater than the conditional indirect utility function for the status quo. The  $i^{th}$  household will be willing to accept the initial bid when  $U_i^1 \geq U_i^0$ .

Therefore, the choice problem can be modeled as binary response variable  $Y$ ;

$$\text{Where, } Y_i = \begin{cases} 1, \text{ if } U_i^1(L - bid, Z, q^*) + e_i \geq U_i^0(L - bid, Z, q) + e_0 \\ 0, \text{ otherwise} \end{cases}$$

Following Hanemann (1984), the probit model can be specified as:  $Y^* = \beta'X_i + \varepsilon_i$   
 $Y_i = 1 \text{ if } Y^* \geq bid1 \text{ and } Y_i = 0 \text{ if } Y^* < bi$

Where:  $\beta$ =is vector of unknown parameters of the model,  $X_i$ =is vector of explanatory variables

$Y^*$  = Unobservable households' actual WTP for soil conservation,  $Y_i$  = Discrete response of the respondents for the WTP,  $bid1$  = the offered initial bids assigned arbitrarily to the  $i^{th}$  respondents and  $\varepsilon_i$  = Unobservable random component distributed  $N(0, \sigma)$

### **Estimation of the Mean Willingness to Pay**

The most general econometric model for the double-bounded data is:  $WTP_{ij} = \mu_j + \varepsilon_{ji}$

Where  $WTP_{ij}$  represents the  $i^{th}$  respondent's willingness to pay, and  $j=1,2$  represents the first and second answer. The mean for the first and second responses are represented by  $\mu_1$  and  $\mu_2$ .

Following, Greene, (1997), a Bivariate Probit Model can be specified as:

$$\begin{aligned} y_1^* &= \beta_1 X_1' + \varepsilon_1 \text{ and } y_2^* = \beta_2 X_2' + \varepsilon_2 \\ E(\varepsilon_1/X_1, X_2) &= E(\varepsilon_2/X_1, X_2) = 0 \\ \text{var}(\varepsilon_1/X_1, X_2) &= \text{var}(\varepsilon_2/X_1, X_2) = 1 \\ \text{cov}(\varepsilon_1, \varepsilon_2/X_1, X_2) &= \rho \end{aligned}$$

Where:  $y_1^*$  is  $i^{th}$  respondents' unobservable true WTP at the time of the first bid.  $WTP = 1$  if  $y_1^* > bid1$ ; otherwise zero.  $y_2^*$ : is the  $i^{th}$  respondent implicit underlying point estimate at the time of the second bid is offered. The mean willingness to pay (MWTP) from bivariate probit model can be calculated using the formula specified by Haab and McConnell (2002).  $MWTP = \frac{-\alpha}{\beta}$

Where  $\alpha$ =a coefficient for the constant term, and  $\beta$ =a coefficient offered bids to the respondents

## **Results and Discussion**

### **Descriptive Analysis**

Randomly collected data from 184 respondents were utilized in the analysis. From total sample households 162 (88%) were willing to take and contribute to the pre specified initial offered bid and 22 (12%) of the households were not willing to pay the initially offered pre specified bid.

**Table 3.1: Descriptive statistics of variables by their willingness to pay**

Variable	WTP N=162				NWTP N=22			
	Mean	Std. D.	Min	Max	Mean	Std. D.	Min	Max
<b>Bid1</b>	45.37	22.32	20	80	67.73	23.89	20	80
<b>Ans1</b>	1.00	0.00	1	1	0.00	0.00	0	0
<b>Bid2</b>	90.09	45.58	10	160	33.86	11.95	10	40
<b>Ans2</b>	0.60	0.49	0	1	0.55	0.51	0	1
<b>MaxWTP</b>	89.31	38.50	12	200	38.41	27.27	0	72
<b>Sex</b>	0.81	0.39	0	1	0.64	0.49	0	1
<b>Mstatus</b>	0.93	0.25	0	1	0.91	0.29	0	1
<b>Edu</b>	0.42	0.50	0	1	0.05	0.21	0	1
<b>Fmlysize</b>	7.02	2.30	2	12	6.95	3.00	1	11
<b>Dratio</b>	89.63	62.81	0	300	97.28	39.78	25	166.67
<b>Frmsize</b>	1.24	0.95	0	4.75	1.19	0.86	0.25	3
<b>TLU</b>	3.61	2.82	0.2	16	2.20	1.55	0.7	8.2
<b>Income('000)</b>	20.67	11.85	3.40	62.85	16.92	7.307	3.20	32.60
<b>OffFarminc</b>	0.15	0.36	0	1	0.14	0.35	0	1
<b>SlopeLand</b>	0.66	0.48	0	1	0.50	0.51	0	1
<b>SFerosion</b>	0.86	0.34	0	1	0.77	0.43	0	1
<b>Distmarket</b>	129.11	90.24	10	345	111.59	90.59	15	300
<b>Distcomland</b>	24.14	34.85	1	270	35.68	32.55	1	120
<b>Perc</b>	0.98	0.14	0	1	0.68	0.48	0	1
<b>Attit</b>	0.77	0.43	1	1	0.54	0.51	0	0
<b>FreqDA</b>	23.17	29.85	0	144	12.41	26.79	0	96
<b>ACredit</b>	0.49	0.50	0	1	0.41	0.50	0	1
<b>Assist</b>	0.10	0.30	0	1	0.05	0.21	0	1

The ratio of male WTP percentage to NWTP percentage is larger than the ratio of female WTP percentage to female NWTP percentage. This indicates that relatively male headed household is more WTP than that of female headed household. 98% of the respondents who were willing to take the pre specified initial bid in soil conservation practices perceived soil erosion as a problem in their area. This shows that household head who perceived soil erosion as problem more WTP for soil conservation than those who didn't perceived.

### *Econometrics Analysis of Households' WTP*

Before running the econometric model, the independent variables were tested for the presence of multicollinearity. The result showed that there were no multicollinearity problems between the variables. The value for Contingency Coefficient(CC) for the dummy variables were less than 0.75 and the value of Variance Inflation Factor (VIF) for the continuous variables were less than 10; which indicate that multicollinearity is not a serious problem. The robust standard errors are used to account for any form heteroscedasticity. The Wald chi-square test is used as the measure of overall significance of a model in probit model estimation. The result of the probit model shows that, the probability of the chi-square distributions (73.73) at 19 degrees of freedom is 0.0000, which is significant at 1%. This shows that we reject the null hypothesis which stated that the coefficients of all explanatory variables included in the model are zero.

Table 3.2: Probit model estimation of the WTP for SWC.

Exp. Variables					Marginal Effect			
	Coef.	Std. Err.	z	P>z	dy/dx	Std. Err.	z	P>z
Bid1	-0.04	0.01	-3.69	0.000	0.00	0.001	-4.34	0.000
Sex	0.95	0.39	2.43	0.015	0.08	0.033	2.47	0.014
Mstatus	0.58	0.67	0.86	0.388	0.05	0.056	0.88	0.378
Edu	1.50	0.49	3.08	0.002	0.13	0.044	2.92	0.003
Fmlysize	0.06	0.08	0.81	0.417	0.01	0.007	0.80	0.424
Dratio	0.00	0.00	-0.65	0.515	0.00	0.000	-0.64	0.524
Frmsize	0.19	0.24	0.79	0.428	0.02	0.020	0.82	0.411
TLU	0.26	0.11	2.38	0.017	0.02	0.010	2.22	0.027
Lnincome	0.80	0.44	1.83	0.067	0.07	0.037	1.85	0.065
OffFarminc	0.39	0.40	0.96	0.338	0.03	0.035	0.94	0.346
SlopeLand	0.68	0.39	1.73	0.083	0.06	0.031	1.90	0.057
SFerosion	0.51	0.40	1.28	0.200	0.04	0.034	1.28	0.200
Distmarket	0.01	0.00	2.4	0.016	0.00	0.000	2.43	0.015
Distcomland	0.00	0.01	-0.63	0.526	0.00	0.000	-0.65	0.514
Perc	2.24	0.48	4.69	0.000	0.19	0.037	5.13	0.000
Attit	0.47	0.44	1.05	0.292	0.04	0.036	1.12	0.265
FreqDA	0.01	0.01	2.1	0.036	0.00	0.001	1.99	0.047
ACredit	0.33	0.51	0.65	0.519	0.03	0.042	0.66	0.507
Assist	0.16	0.55	0.29	0.771	0.01	0.048	0.29	0.772
_cons	-11.05	4.66	-2.37	0.018				
Num. Obs.		184				184		

Factors like *Sex*, education, TLU, income, land slope, market distance, perception, and frequency of extension agent have positive and significant influence on the probability of farmers' WTP for soil conservation. As sex is male the probability of yes for WTP is increase by 8%. Female headed households have fewer resources needed for conservation than male headed households. The marginal effect shows that the probability of the WTP for soil conservation increases by 13 % if the household head is literate. It is clear that education increases environmental awareness and enhance their understanding of the value for environmental goods. The probability of WTP increases by 2% for a unit increase in TLU. Therefore, having a large number of livestock can strengthen farmers' WTP because of they are highly affected by the degradation of the grazing land used for their livestock. The effect of farmer's perception on WTP is positive and statistically significant at 1%. Additional contact with the extension agents increases the probability of farmers' willingness to pay for soil conservation practices by 0.11%.

The findings show that those farmers with steep land have a positive and significant effect on the probability of their WTP. Farmers who are far away from the market have a positive and significant influence on their WTP for soil conservation which is unexpected result. This may be because of low opportunity cost of labor contribution (as a result of a possible absence of adequate labor market for alternative employment) as a farmer far away from the market center. Bid one is a significant explanatory variable in this model. It is statistically significant at 1% level. The coefficient shows that as the price offered to a household increases the farmers' demand for soil conservation of communal land will decrease, keeping other factor constant.

**Bivariate Probit Model**

The test against the null hypothesis is that the two equations can be independently estimated or the correlation between the two error terms is Zero can be checked by looking at the Likelihood-ratio test of the correlation coefficient (i.e.  $\rho=0$ ). The result is significant at 1 % level of significance, showing that the two equations can be estimated simultaneously.

**Table 3.3: Bivariate Probit Statistical Regression output**

Robust						
Variable	Coef.	Std. Err.	z	P>z	[95% Conf.	Interval]
Ans1						
Bid1	-0.0337	0.0058	-3.57	0.000	-0.0604278	-0.0232
Ans2						
Bid2	-0.0148	0.0023	-6.34	0.000	-0.0240276	-0.0126043

**Marginal effect after biprobit**

Variable	dy/dx	Std. Err.	z	P>z	[95% C.I.]	X
Bid1	-.00303	.00007	-0.45	0.005	-.000162 .000101	48.3696
Bid2	-.0077792	.00128	-6.06	0.000	-.010296 -.005263	83.3696

The coefficients of both the initial and second bids are statistically significant at 1%. The coefficients show that as the price offered to a household increases the farmers demand for soil conservation decreases keeping other factor constant. The marginal effect of full model for both Bid1 and Bid2 is significant at 1% level of significance. As Bid1 increase by one unit, WTP decrease by 0.03%; and as Bid2 increase by 1 unit WTP decrease by 0.07%.

**Summary of Households' WTP and Estimation of Total Welfare  
Estimation of Mean from Double Bounded Dichotomous Format**

**Table 3.4: Descriptive Statistics of the Dichotomous Choice Format**

Variable	Obs	Mean	Stan. Dev	Min	Max
Bid1	184	48.369	23.622	20	80
Bid2	184	83.342	46.710	10	160
YBid1	184	0.880	0.325	0	1
YBid2	184	0.598	0.492	0	1

Source: Own Survey, 2014

The result shows that the average initial bid was 48.37 Labor days per year. Whereas, the average second bid was 83.34 Labor days per year for soil conservation. The “yes” response for the first bid is 88% while 59.8% for the second bid. Both results imply that higher initial bid and second bid lead to lower probability of accepting that bid.

**Table 3.5: Estimates of the Double Bounded Dichotomous Choice Format.**

	Robust			
	Coef.	S.E.	Z	P>z
Bid1	-0.0336697	0.0058	-3.57	0.000
Constant	2.314392	0.3776	6.13	0.000
Bid2	-0.0147951	0.0023	-6.34	0.000
Constant	1.508658	0.2514	6.00	0.000
/athrho	0.752	0.324	2.32	0.020
Rho	0.636	0.1928		
Wald test of rho=0: chi2(1)=5.38423; Prob > chi2 = 0.0203				
Log pseudolikelihood = -166.38832				
Number of obs = 184				
Wald chi2(2) = 47.06				
Prob > chi2 = 0.0000				

Using the coefficients of bid and constant in Table 3.5 mean WTP for soil conservation practices from the double bounded probit estimate was estimated using the formula by Haab and McConnell (2002) to be 85.36 Labor days per year per household. At 95% confidence interval the WTP varies between 68.68 to 101.94 Labor days per year.

### *Welfare Measures and Aggregation*

Aggregation of benefit (TEV) is the final step in the CV research. Measurement of welfare using WTP mean is very important in aggregation of benefit (Alemu, 2000)

**Table 3.6: Welfare Measures and Aggregate Benefits by Peasant Associations**

Name of kebeles	Total HHs of Kebeles	Sampled HHs with valid response	Sampled HHs with invalid response	Proportion of invalid response	Expected Total HH with invalid response	Expected Total HH with valid response	Mean WTP	Total WTP
H/Cari	3150	40	0	0.000	000.00	3150.00	85.36	268884.00
R/kolati	2110	37	3	0.075	158.25	1951.75	85.36	166601.38
J/kerensa	2550	34	6	0.15	382.5	2167.5	85.36	185017.80
W/gose	3000	38	2	0.05	150	2850	85.36	243276.00
D/kerensa	2400	35	5	0.125	300	2100	85.36	179256.00
<b>Total</b>	<b>13,210</b>	<b>184</b>	<b>16</b>		<b>990.75</b>	<b>12219.25</b>		<b>1,043,035</b>

In Table 3.6 the aggregate WTP in terms of labor days per year was converted into monetary value by using wage rate in the local labor market. Shadow wage rate which is good measure of economic value didn't available. Therefore, based on the mean from the double bounded dichotomous choice format, the aggregate WTP for soil conservation practices is 1,043,035 Labor days per year (18,774,630 ETB) in the selected five kebeles. This shows that there is high level of WTP for soil conservation practices in the study area. The total willingness to pay in the whole study area is simply the product of the respective means and the number of expected households who have valid responses. Hence, the aggregate value of soil conservation in the study area from the double bounded are 10,909,349 person days per annum (196,368,282 ETB).

### **Conclusion**

Both descriptive and econometric models were employed for the analysis. The result of descriptive statistics indicated that most of the respondents have perceived the problem of soil erosion and are willing to pay for conservation practices. The econometric result also show that sex of household head, education level of the head, amount of livestock, and income



have a positive and significant influence on the probability of willingness to pay for soil conservation. Variables such as slope of land, perception about soil erosion problem, and the frequency of extension visit have also a positive and significant effect on the WTP for soil conservation. The study also shows that the mean willingness to pay estimated from the Double Bounded Dichotomous Choice was computed at 85.36 person days per annum. The total aggregate value of soil conservation is 10,909,349 person days per annum (196,368,282 ETB) from double bound CVM. Therefore, this study concludes that farmers have perceived the problem of soil erosion and to overcome the problem, they are willing to pay in terms of labor contribution for improved soil conservation practice.

### ***Policy Implications***

The findings of this study indicate that any plan for intervention in the soil conservation should recognize the existing heterogeneity in household characteristics, perception on soil-erosion, physical factors, institutional and demographic factors. In this respect, increasing the level of farmers' perception on soil conservation through training and centered participatory will help them to develop a positive attitude towards conservation work. Moreover, policy makers can target areas where soil erosion is severe so as to plan more appropriate projects in the future. Policy makers should take into account that farmers who have educated have higher probability of accepting soil conservation projects than those who are not educated. The interaction between extension agents and farmers should be strengthened. More frequent contact between the two could enhance farmers understanding of the environmental problems in general and soil erosion problems in particular. The relatively better off farmers are more willing to contribute to the soil conservation efforts in the study area. Policy makers may target these households especially at the beginning of the conservation efforts based on their willingness to pay.

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